

例7.3. 假设 \mathbf{X} 与 \mathbf{Y} 相互独立, $\mathbf{X} \sim \mathcal{P}(\lambda_1)$, $\mathbf{Y} \sim \mathcal{P}(\lambda_2)$.

令 $\mathbf{Z} = \mathbf{X} + \mathbf{Y}$. 求: \mathbf{Z} 的分布; 在 $\mathbf{Z} = n$ 的条件下, \mathbf{X} 的条件分布.

- $\mathbf{Z} \sim \mathcal{P}(\lambda_1 + \lambda_2)$: $\forall n \geq 0, \quad \mathbf{P}(\mathbf{Z} = n)$

$$= \sum_{k=0}^n \mathbf{P}(\mathbf{X} = k, \mathbf{Y} = n - k) = \sum_{k=0}^n \frac{\lambda_1^k}{k!} e^{-\lambda_1} \frac{\lambda_2^{n-k}}{(n-k)!} e^{-\lambda_2}$$

$$= \frac{1}{n!} \left(\sum_{k=0}^n \binom{n}{k} \lambda_1^k \lambda_2^{n-k} \right) e^{-(\lambda_1 + \lambda_2)} = \frac{(\lambda_1 + \lambda_2)^n}{n!} e^{-(\lambda_1 + \lambda_2)}$$

- $n > 0$ 时, 记 $p = \frac{\lambda_1}{\lambda_1 + \lambda_2}$, $q = 1 - p$. 则, $k = 0, 1, \dots, n$,

$$\mathbf{P}(\mathbf{X} = k | \mathbf{Z} = n) = \frac{\binom{n}{k} p^k q^{n-k}}{(\lambda_1 + \lambda_2)^n} = \mathbf{C}_n^k p^k q^{n-k},$$

- 条件分布是二项分布.

例7.4. 假设 \mathbf{X} 与 \mathbf{Y} 相互独立, $\mathbf{X} \sim \mathbf{B}(n_1, p)$, $\mathbf{Y} \sim \mathbf{B}(n_2, p)$, $0 < p < 1$. 令 $\mathbf{Z} = \mathbf{X} + \mathbf{Y}$. 求: \mathbf{Z} 的分布; 在 $\mathbf{Z} = n$ 的条件下, \mathbf{X} 的条件分布.

- $n = 0, 1, \dots, n_1 + n_2$. 记 $q = 1 - p$, 则 $P(\mathbf{Z} = n)$

$$= \sum_{k=0}^n P(\mathbf{X} = k, \mathbf{Y} = n - k) = \sum_{k=0}^n C_{n_1}^k C_{n_2}^{n-k} p^{k+n-k} q^{n_1-k+n_2-(n-k)}$$

$$= \sum_{k=0}^n C_{n_1}^k C_{n_2}^{n-k} p^n q^{n_1+n_2-n}$$

$$= C_{n_1+n_2}^n p^n q^{n_1+n_2-n}$$

- $n = 0$ 时, $P(\mathbf{X} = 0 | \mathbf{Z} = 0) = 1$.

- $n > 0$ 时, $k = 0, 1, \dots, n$,

$$P(\mathbf{X} = k | \mathbf{Z} = n) = \frac{C_{n_1}^k C_{n_2}^{n-k}}{C_{n_1+n_2}^n}$$

- 条件分布是超几何分布.

- 定理4.1. 设 (X, Y) 有联合密度 $p(x, y)$, $Z = X + Y$. 则

$$p_Z(z) = \int_{-\infty}^{\infty} p(x, z - x) dx.$$

- 证: 第一步,

$$F_Z(z) = P(X + Y \leq z) = \iint_{x+y \leq z} p(x, y) dx dy.$$

第二步

$$\int_{-\infty}^{\infty} \left(\int_{-\infty}^{z-x} p(x, y) dy \right) dx = \int_{-\infty}^z \int_{-\infty}^{\infty} p(x, u-x) dx du$$

- 推论(系4.1): 若 X, Y 相互独立, 分别有密度 p_X, p_Y , 则 $Z = X + Y$ 是连续型, 且

$$p_Z(z) = \int_{-\infty}^{\infty} p_X(x) p_Y(z - x) dx = \int_{-\infty}^{\infty} p_X(z - y) p_Y(y) dy.$$

例4.1 & 4.2. 设 (\mathbf{X}, \mathbf{Y}) 服从二维正态分布, 联合密度 $p(\mathbf{x}, \mathbf{y})$ 为

$$p(\mathbf{x}, \mathbf{y}) = \hat{\mathbf{C}} \exp \left\{ -\frac{\mathbf{u}^2 - 2 \mathbf{u} \mathbf{v} + \mathbf{v}^2}{2(1 - \rho^2)} \right\}, \quad \left(\mathbf{u} = \frac{\mathbf{x} - \mu_1}{1}, \mathbf{v} = \frac{\mathbf{y} - \mu_2}{2} \right),$$

其中, $\hat{\mathbf{C}} = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}}$. 求 $\mathbf{Z} = \mathbf{X} + \mathbf{Y}$ 的密度.

• $p_Z(z) = \int_{-\infty}^{\infty} p(\mathbf{x}, z - \mathbf{x}) d\mathbf{x}$. 当 \mathbf{y} 取 $z - \mathbf{x}$ 时,

其中, $\mathbf{C} = (z - \mu_1 - \mu_2) / 2$

此时, $\mathbf{u}^2 - 2 \mathbf{u} \mathbf{v} + \mathbf{v}^2$

$$\begin{aligned} &= \mathbf{u}^2 - 2 \mathbf{u} \left(\mathbf{C} - \frac{1\mathbf{u}}{2} \right) + \left(\mathbf{C} - \frac{1\mathbf{u}}{2} \right)^2 \\ &= \left(1 + 2 \frac{1}{2} + \left(\frac{1}{2} \right)^2 \right) \mathbf{u}^2 - 2 \left(\mathbf{C} + \frac{1}{2} \right) \mathbf{C} \mathbf{u} + \mathbf{C}^2. \end{aligned}$$

- 目标: 计算 $p_Z(z) = \int_{-\infty}^{\infty} p(x, z-x) dx$. 已有:

$$p(x, z-x) = \hat{C} \left\{ -\frac{Au^2 - 2Bu + C^2}{2(1-\rho^2)} \right\}, \quad \text{其中, } u = \frac{x - \mu_1}{\sigma_1}$$

$$A = 1 + 2 \frac{\sigma_2^2}{\sigma_1^2} + \left(\frac{\sigma_2}{\sigma_1} \right)^2, \quad B = \left(\rho + \frac{\sigma_2}{\sigma_1} \right) C, \quad C = \frac{z - (\mu_1 + \rho \mu_2)}{\sigma_2}$$

配方:

$$Au^2 - 2Bu + C^2 = A \left(u - \frac{B}{A} \right)^2 - \left(\frac{B^2}{A} - C^2 \right)$$

$$= \hat{C} \exp \left\{ \frac{\frac{B^2}{A} - C^2}{2(1-\rho^2)} \right\} \times \int_{-\infty}^{\infty} \exp \left\{ -\frac{A \left(u - \frac{B}{A} \right)^2}{2(1-\rho^2)} \right\} du$$

$$= \tilde{C} \exp \left\{ \frac{B^2 - AC^2}{2(1-\rho^2)A} \right\}. \quad \tilde{C} = \hat{C} \sqrt{2 \frac{1-\rho^2}{A}} = \frac{1}{\sqrt{2} \frac{\sigma_2}{\sigma_1} A}$$

- 已有: $\mathbf{p}_Z(\mathbf{z}) = \tilde{\mathbf{C}} \exp \left\{ \frac{B^2 - AC^2}{2(1-\rho^2)A} \right\}$, 其中 $\tilde{\mathbf{C}}$ 是常数,

$$\mathbf{A} = 1 + 2 \frac{1}{2} + \left(\frac{1}{2} \right)^2, \quad \mathbf{B} = \left(+ \frac{1}{2} \right) \mathbf{C}, \quad \mathbf{C} = \frac{\mathbf{z} - (\boldsymbol{\mu}_1 + \boldsymbol{\mu}_2)}{2}.$$

- $B^2 - AC^2$

因此,

$$= \left(\left(+ \frac{1}{2} \right)^2 - \mathbf{A} \right) \mathbf{C}^2 = (2 - 1) \frac{(\mathbf{z} - (\boldsymbol{\mu}_1 + \boldsymbol{\mu}_2))^2}{2}.$$

$$\mathbf{p}_Z(\mathbf{z}) = \tilde{\mathbf{C}} \exp \left\{ \frac{(\mathbf{z} - (\boldsymbol{\mu}_1 + \boldsymbol{\mu}_2))^2}{2} \right\}.$$

其中, $\boldsymbol{\mu} = \boldsymbol{\mu}_1 + \boldsymbol{\mu}_2$, $2 = \frac{2}{2} \mathbf{A} = \frac{2}{1} + \frac{2}{2} + 2 \frac{1}{2}$.

- 特别地, 若 $\rho = 0$ (即 \mathbf{X}, \mathbf{Y} 相互独立), 则

$$\mathbf{X} + \mathbf{Y} \sim \mathbf{N}(\boldsymbol{\mu}_1 + \boldsymbol{\mu}_2, \frac{2}{1} + \frac{2}{2}).$$

- 定理4.2. 设 (\mathbf{X}, \mathbf{Y}) 有联合密度 $p(\mathbf{x}, \mathbf{y})$.

令 $\mathbf{Z} = \mathbf{X}/\mathbf{Y}$ (当 $\mathbf{Y} = 0$ 时, 规定 $\mathbf{Z} = 0$). 则 \mathbf{Z} 为连续型, 且

$$p_Z(\mathbf{z}) = \int_{-\infty}^{\infty} |y|p(\mathbf{zy}, \mathbf{y})d\mathbf{y}.$$

- 证明: 第一步, $\frac{x}{y} \leq z$ 当且仅当 “ $y > 0$ 且 $x \leq yz$ ” 或者 “ $y < 0$ 且 $x \geq yz$.” 于是,

$$F_Z(z) = P(Y > 0, X \leq Yz) + P(Y < 0, X \geq Yz).$$

$$= P(Y > 0, X \leq Yz) + P(Y < 0, X \geq Yz)$$

$$= \int_0^{\infty} \int_{-\infty}^{yz} p(\mathbf{x}, \mathbf{y})d\mathbf{x}d\mathbf{y} = \int_0^{\infty} \int_{-\infty}^z p(\mathbf{yu}, \mathbf{y})\mathbf{y}d\mathbf{u}d\mathbf{y}$$

$$= \int_{-\infty}^z \left(\int_0^{\infty} \mathbf{y}p(\mathbf{yu}, \mathbf{y})d\mathbf{y} \right) d\mathbf{u}.$$

- 类似处理**, 即可.

例4.4. \mathbf{X}, \mathbf{Y} 相互独立, 都服从 $\mathbf{N}(0, 1)$. 求 $\mathbf{Z} = \mathbf{X}/\mathbf{Y}$ 的密度.

● 联合密度:

$$p(\mathbf{x}, \mathbf{y}) = \frac{1}{2} \exp \left\{ -\frac{\mathbf{x}^2 + \mathbf{y}^2}{2} \right\}.$$

● 因此,

$$\begin{aligned}
 p_Z(z) &= \int_{-\infty}^{\infty} |y| p(z y, y) dy = \int_{-\infty}^{\infty} |y| \frac{1}{2} \exp \left\{ -\frac{(z y)^2 + y^2}{2} \right\} dy \\
 &= \frac{1}{2} \int_0^{\infty} e^{-(z^2+1)u} du = \frac{1}{(z^2+1)}.
 \end{aligned}$$

- 定理4.3. 假设 (X, Y) 为连续型, 有密度 $p(x, y)$.

假设

$$(U, V), \quad \text{其中 } U = f(X, Y), \quad V = g(X, Y).$$

如果(1) $P((U, V) \in A) = 1$ 且 $(f, g) : A \rightarrow G$ 是一对一的;

(2) $f, g \in C^1(A)$, 且 $\frac{\partial(u, v)}{\partial(x, y)} \neq 0, \forall (x, y) \in A$,

那么, (U, V) 是连续型, 且

$$p_{U, V}(u, v) = p(x(u, v), y(u, v)) \left| \frac{\partial(x, y)}{\partial(u, v)} \right|, \quad (u, v) \in G.$$

- 证: $\forall D \subseteq G$, 找 $D^* \subseteq A$ 使得 $(U, V) \in D$ iff $(X, Y) \in D^*$. 于是,

$$P(D) = P(D^*) = \iint_{D^*} p(x, y) dx dy = \iint_D p_{U, V}(u, v) du dv.$$

例4.5, 4.7, & 习题三、21. 假设 \mathbf{X}, \mathbf{Y} 相互独立, 都服从 $\mathbf{N}(0, 1)$.

- 用极坐标表达:

$$\mathbf{X} = \mathbf{R} \cos \Theta, \quad \mathbf{Y} = \mathbf{R} \sin \Theta.$$

- $\mathbf{A} = \{(\mathbf{x}, \mathbf{y}) : \mathbf{x} \neq 0, \mathbf{y} \neq 0\}$,

$$\mathbf{G} = \left\{ \begin{pmatrix} r \\ \theta \end{pmatrix} : r > 0, 0 < \theta < 2\pi; \theta \neq \frac{\pi}{2}, \frac{3\pi}{2} \right\}.$$

$$\frac{\partial(x, y)}{\partial(r, \theta)}$$

$$\cos \theta \quad r \sin \theta$$

$$\sin \theta$$

- $p_{R, \Theta}(r, \theta) = \frac{1}{2\pi} e^{-\frac{1}{2}(x^2+y^2)} \left| \frac{\partial(x, y)}{\partial(r, \theta)} \right|$
 $= \frac{1}{2\pi} r e^{-\frac{1}{2}r^2}, \quad r > 0, 0 < \theta < 2\pi; \theta \neq \frac{\pi}{2}, \frac{3\pi}{2}.$
- \mathbf{R}, Θ 独立: $p_{R, \Theta}(r, \theta) = p_{\mathbf{R}}(r) \cdot p_{\Theta}(\theta).$

- $W := R^2 = X^2 + Y^2 \sim \text{Exp}(\frac{1}{2})$. 因为, $\forall w > 0$,

$$p_W(w) = p_R(r) \frac{dr}{dw} = r \exp\left\{-\frac{r^2}{2}\right\} \cdot \frac{1}{2r} = \frac{1}{2} e^{-\frac{w}{2}}.$$

- $U := e^{-\frac{1}{2}W} \sim U(0, 1)$. 因为, $\forall p \in (0, 1)$,

$$P(U \leq p) = P(W \geq -2 \ln p) = e^{-\frac{1}{2}(-2 \ln p)} = e^{\ln p} = p.$$

- $V = \Theta \sim U(0, 1)$, U, V 相互独立.

- $R = \sqrt{-2 \ln U}$, $\Theta = 2 V$, 即

$$X = \sqrt{-2 \ln U} \cos(2 V), \quad Y = \sqrt{-2 \ln U} \sin(2 V).$$

2. 两个随机变量的函数的数学期望

- 随机向量函数的期望(定理4.6):

离散型: $E\mathbf{f}(\mathbf{X}, \mathbf{Y}) = \sum_{i,j} \mathbf{f}(x_i, y_j) P(\mathbf{X} = x_i, \mathbf{Y} = y_j).$

连续型: $E\mathbf{f}(\mathbf{X}, \mathbf{Y}) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \mathbf{f}(x, y) p(x, y) dx dy.$

- 定理4.4. 若 \mathbf{X} 与 \mathbf{Y} 相互独立, 则 $E\mathbf{X}\mathbf{Y} = (E\mathbf{X})(E\mathbf{Y}).$

- 例, 连续型的证明:

$$E\mathbf{X}\mathbf{Y} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xyp_X(x)p_Y(y) dx dy = (E\mathbf{X})(E\mathbf{Y}).$$

- 定理4.5. 若 \mathbf{X} 与 \mathbf{Y} 相互独立,

则 $\text{var}(\mathbf{X} + \mathbf{Y}) = \text{var}(\mathbf{X}) + \text{var}(\mathbf{Y}).$

- 证: 左 = $E(\mathbf{X} + \mathbf{Y} - (E\mathbf{X} + E\mathbf{Y}))^2$
= 右 + $2E(\mathbf{X} - E\mathbf{X})(\mathbf{Y} - E\mathbf{Y}).$