

Supplementary Material for “Online Estimation for Functional Data”

In this Supplementary Material, we first derive the asymptotic distributions of the proposed mean and covariance estimates, i.e., Theorem 1 and 2, in Section S.1 and S.2, respectively. The convergence of bandwidth selection and lower bound of relative efficiencies are shown in Section S.3 and S.4.

S.1. PROOF OF THEOREM 1

Proof. This proof is in analogy to the proof of Theorem 3.1 in Zhang and Wang (2016). Denote the current time by K and recall $e_{:,k}^{(K)} :; k = 1; \dots; K$ is the pseudo-bandwidth chain at time K . We omit the superscript “ (K) ” of $e_{:,k}^{(K)}$, $\hat{A}_{:,j}^{(K)}$ in (18), $\hat{\gamma}_{:,j}^{(K)}$ in (14) and $e^{(K)}$ in (10) in the proof when no confusion exists. The online estimate can be written as

$$e(t) = \frac{R_0 Q_2}{Q_0 Q_2} \frac{R_1 Q_1}{Q_1^2}$$

where for $r = 0; 1; 2$,

$$\begin{aligned} Q_r &= \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} W_{:,k}(T_{kij} - t) \frac{T_{kij}}{e_{:,k}} t^{-r} = \sum_k !_k Q_{r,k}; \\ R_r &= \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} W_{:,k}(T_{kij} - t) \frac{T_{kij}}{e_{:,k}} t^{-r} Y_{kij} = \sum_k !_k R_{r,k}; \end{aligned}$$

We further denote for $r = 0; 1$,

$$\mathcal{G}_r = \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} W_{:,k}(T_{kij} - t) T_{kij}^r = \sum_k !_k \mathcal{G}_{r,k};$$

and define

$$e'(t) = \frac{1}{e_{:,k}} \frac{R_1}{Q_2} \frac{R_0 Q_1 - Q_0}{Q_1^2 - Q_0};$$

Note that

$$Q_1 = \sum_k !_k Q_{1,k} = \sum_k !_k \frac{\mathcal{G}_{1,k} - t \mathcal{G}_{0,k}}{e_{:,k}};$$

one can derive

$$e(t) = \frac{R_0}{Q_0} \sum_k !_k \frac{\mathcal{G}_{1,k}}{Q_0} e'_k(t) + t \sum_k !_k \frac{\mathcal{G}_{0,k}}{Q_0} e'_k(t);$$

We also define

$$\bar{e}(t) = \frac{R_0}{Q_0} \times \sum_k \frac{\mathbb{E}_{1:k}}{Q_0} '(t) + t \sum_k \frac{\mathbb{E}_{0:k}}{Q_0} '(t); \quad (\text{S.1})$$

where $'(\cdot)$ is the first derivative of (\cdot) . Then

$$e(t) - \bar{e}(t) = \sum_k \frac{\mathbb{E}_{1:k}}{Q_0} \frac{Q_0(R_1 - Q_1 - e_{:k}'Q_2)}{Q_2Q_0} - \frac{Q_1(R_0 - Q_0 - e_{:k}'Q_1)}{Q_1^2}; \quad (\text{S.2})$$

It is straightforward to show that both Q_0 and $Q_0Q_2 - Q_1^2$ are positive and bounded away from 0 with probability tending to one,

$$Q_1 = O_p \left(\beta_1 + \frac{\beta_1^{-1}}{S_{K,1}} \right)^{\frac{1}{2}}; \\ (R_1 - Q_1 - e_{:k}'Q_2)(R_0 - Q_0 - e_{:k}'Q_1) = O_p \left(\beta_2 + \beta_1 e_{:k} + \frac{\beta_1^{-1}}{S_{K,1}} + \frac{S_{K,2}}{S_{K,1}^2} \right)^{\frac{1}{2}} A.$$

Then by the Cramer-Wold device and Lyapunov condition, we can achieve the asymptotic joint normality of $(R_0 - ER_0; \mathbb{E}\mathbb{G}_1 - E\mathbb{G}_1; \mathbb{E}\mathbb{G}_0 - E\mathbb{G}_0)$, where the rate of convergence is $\min[(\beta_1^{-1} = S_{K,1})^{1/2}; fQ_2 = S_{K,1}^2g^{1/2}]$.

Explicitly for $r, r' = 0, 1$,

$$\begin{aligned} \mathbb{E}(\mathbb{G}_r) &= t^r f(t) + \frac{1}{2} (\mathcal{W}) f' t^r f'(t) + t^r f''(t) g_{:2} + o_p(\beta_2); \\ \mathbb{E}(R_0) &= f(t) f(t) + \frac{1}{2} (\mathcal{W}) f'(t) f''(t) + 2 f'(t) f'(t) + f'' f(t) g_{:2} + o_p(\beta_2); \\ \text{Cov}(\mathbb{G}_r, \mathbb{G}_{r'}) &= \frac{\beta_1^{-1}}{S_{K,1}} R(\mathcal{W}) t^{r+r'} f(t) + o_p \left(\frac{\beta_1^{-1}}{S_{K,1}} \right); \\ \text{Var}(R_0) &= \frac{\beta_1^{-1}}{S_{K,1}} R(\mathcal{W}) (f^2(t) + f'(t)f(t) + f''(t)f(t)) f(t) + \frac{S_{K,2}}{S_{K,1}^2} (t^2 f(t)^2) + o_p \left(\frac{\beta_1^{-1}}{S_{K,1}} + \frac{S_{K,2}}{S_{K,1}^2} \right); \\ \text{Cov}(R_0, \mathbb{G}_r) &= \frac{\beta_1^{-1}}{S_{K,1}} R(\mathcal{W}) t^r f(t) f(t) + o_p \left(\frac{\beta_1^{-1}}{S_{K,1}} \right); \end{aligned}$$

From the delta method, $\bar{e}(t)$ follows the asymptotic normality that

$$\bar{e}(t) - \bar{e}(t) = \frac{1}{2} (\mathcal{W}) f''(t) g_{:2} + o_p(\beta_2) \stackrel{d}{\rightarrow} N(0; 1);$$

where

$$= R(\mathcal{W}) \frac{(t; t) + f'^2(t)}{S_{K,1} f(t)} \beta_1 + (t; t) \frac{S_{K,2}}{S_{K,1}^2}.$$

When $\bar{m}_{:K} = (N_K)^{1/4} / 0$ and $\beta_1 = S_{K,1}^{-1/5}$, the first term in above dominates. When $\bar{m}_{:K} = (N_K)^{1/4} / C$ and $\beta_1 = S_{K,1}^{-1/5}$, where $0 < C < 1$, the two terms of are of the same

order as θ^4 . When $m_{,K} = (N_K)^{1/4} \neq 1$ and $\theta = S_{K,1}^{-1/5}$, the second term of $\hat{\theta}$ dominates and the bias becomes negligible. Noting (S.2), the proof of Theorem 1 is completed. ■

S.2. PROOF OF THEOREM 2 AND LEMMA 1

We omit the superscript " (K) " of $e_{,k}^{(K)}$, $\theta_j^{(K)}$ and $e^{(K)}$ in this section when no confusion exists. The proof of Lemma 1 is similar to the proof of Theorem 2, with $j:j \geq Z$ replaced by θ . Before delineating the proof of Theorem 2, we state the following lemma which is a preparation for the proof of Theorem 2.

Lemma S. 1. *Let*

$$Q_{0R_1} = \frac{1}{S_{K,1}} \sum_{k'=1}^K \sum_{i'=1}^{K'} \sum_{j'=1}^{K''} W_{-,k'}(T_{k'i'j'}) - T_{kil}) ;$$

$$\bar{R}_1 = S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^k \sum_{j \neq i} \sum_{k'=1}^K \sum_{i'=1}^{K'} \sum_{j'=1}^{K''} W_{-,k'}(T_{kij} - s) W_{-,k'}(T_{kil} - t) W_{-,k'}(T_{k'i'j'} - T_{kil}) ;$$

$$(Y_{kij} - (T_{kij})) (Y_{k'i'j'} - (T_{k'i'j'})) - \frac{1}{2}''(T_{kil})(T_{k'i'j'} - T_{kil})^2 ;$$

then

$$E(\bar{R}_1 - Q_{0R_1}) = O_p(S_{K,1}^{-1} S_{K,2}^{-1} S_{K,3}^{-1});$$

$$\text{Var}(\bar{R}_1 - Q_{0R_1}) = O_p(S_{K,1}^{-2} S_{K,2}^{-2} S_{K,3}^2 + \sum_k S_{k,3} e_{,k}^{-1} e_{,k}^{-2} + \sum_k S_{k,4} e_{,k}^{-1} e_{,k}^{-1} + \sum_k S_{k,5} e_{,k}^{-1} + S_{K,2}^{-2} \sum_k S_{k,4} e_{,k}^4) ;$$

Proof. By easy calculation,

$$E \bar{R}_1 = S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^k \sum_{j \neq i} \sum_{k'=1}^K \sum_{i'=1}^{K'} \sum_{j'=1}^{K''} E W_{-,k'}(T_{kij} - s) W_{-,k'}(T_{kil} - t) W_{-,k'}(T_{k'i'j'} - T_{kil}) (Y_{kij} - (T_{kij})) (Y_{k'i'j'} - (T_{k'i'j'}))$$

$$= S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^k \sum_{j \neq i} \sum_{i'=1}^{K'} E W_{-,k}(T_{kij} - s) W_{-,k}(T_{kil} - t) W_{-,k}(T_{kij'} - T_{kil}) (Y_{kij} - (T_{kij})) (Y_{kij'} - (T_{kij'}))$$

$$= S_{K,1}^{-1} S_{K,2}^{-1} \sum_{k=1}^K \sum_j m_{ki}^2 (m_{ki} - 1) E W_{-,k}(T_{kij} - s) W_{-,k}(T_{kil} - t)$$

$$\# \\ \mathcal{W}_{\sim_{\mathcal{K}}}(T_{kij'} - T_{kil})(Y_{kij} - (T_{kij}))(Y_{kij'} - (T_{kij'}))$$

$$= O_p(S_{K,1}^{-1}S_{K,2}^{-1}S_{K,3}) ;$$

and hence $E \mathcal{R}_1 = Q_{0R_1} = O_p(S_{K,1}^{-1}S_{K,2}^{-1}S_{K,3})$. Note that

$$\begin{aligned} \mathcal{R}_1^2 &= S_{K,1}^{-2}S_{K,2}^{-2} \times \\ &\quad W_{k_1 i_1 j_1}(s) W_{k_1 i_1 l_1}(t) W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1}) W_{k_2 i_2 j_2}(s) W_{k_2 i_2 l_2}(t) W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ &\quad (Y_{k_1 i_1 j_1} - k_1 i_1 j_1) (Y_{k'_1 i'_1 j'_1} - k'_1 i'_1 j'_1) \frac{1}{2}''_{k_1 i_1 l_1} (T_{k'_1 i'_1 j'_1} - T_{k_1 i_1 l_1})^2 \\ &\quad (Y_{k_2 i_2 j_2} - k_2 i_2 j_2) (Y_{k'_2 i'_2 j'_2} - k'_2 i'_2 j'_2) \frac{1}{2}''_{k_2 i_2 l_2} (T_{k'_2 i'_2 j'_2} - T_{k_2 i_2 l_2})^2 ; \end{aligned}$$

$E \mathcal{R}_1^2$ is the summation of the following quantities:

$$\begin{aligned} (a) &= S_{K,1}^{-2}S_{K,2}^{-2} E \times \\ &\quad W_{k_1 i_1 j_1}(s) W_{k_1 i_1 l_1}(t) W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1}) W_{k_2 i_2 j_2}(s) W_{k_2 i_2 l_2}(t) W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ &\quad (Y -)_{k_1 i_1 j_1} (Y -)_{k'_1 i'_1 j'_1} (Y -)_{k_2 i_2 j_2} (Y -)_{k'_2 i'_2 j'_2} ; \\ (b) &= \frac{1}{2} S_{K,1}^{-2}S_{K,2}^{-2} E \times \\ &\quad W_{k_1 i_1 j_1}(s) W_{k_1 i_1 l_1}(t) W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1}) W_{k_2 i_2 j_2}(s) W_{k_2 i_2 l_2}(t) W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ &\quad (Y -)_{k_1 i_1 j_1} (Y -)_{k_2 i_2 j_2} ''_{k_2 i_2 l_2} (T_{k'_2 i'_2 j'_2} - T_{k_2 i_2 l_2})^2 ; \\ (c) &= \frac{1}{2} S_{K,1}^{-2}S_{K,2}^{-2} E \times \\ &\quad W_{k_1 i_1 j_1}(s) W_{k_1 i_1 l_1}(t) W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1}) W_{k_2 i_2 j_2}(s) W_{k_2 i_2 l_2}(t) W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ &\quad (Y -)_{k_1 i_1 j_1} (Y -)_{k_2 i_2 j_2} (Y -)_{k'_2 i'_2 j'_2} ''_{k_1 i_1 l_1} (T_{k'_1 i'_1 j'_1} - T_{k_1 i_1 l_1})^2 ; \\ (d) &= \frac{1}{4} S_{K,1}^{-2}S_{K,2}^{-2} E \times \\ &\quad W_{k_1 i_1 j_1}(s) W_{k_1 i_1 l_1}(t) W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1}) W_{k_2 i_2 j_2}(s) W_{k_2 i_2 l_2}(t) W_{k'_2 i'_2 j'_2}(T_{k_2 i_2 l_2}) \\ &\quad (Y -)_{k_1 i_1 j_1} (Y -)_{k_2 i_2 j_2} ''_{k_1 i_1 l_1} (T_{k'_1 i'_1 j'_1} - T_{k_1 i_1 l_1})^2 ''_{k_2 i_2 l_2} (T_{k'_2 i'_2 j'_2} - T_{k_2 i_2 l_2})^2 ; \end{aligned}$$

which are computed as follows.

- a. For (a), there are four cases according to the relationship between $(k_1; i_1)$, $(k_2; i_2)$, $(k'_1; i'_1)$, $(k'_2; i'_2)$ to make the quantity nonzero:

(B.1) Let $\stackrel{a}{1} = f(k_1; i_1) = (k'_1; i'_1); (k_2; i_2) = (k'_2; i'_2); (k_1; i_1) \not\in (k_2; i_2); j_1 \not\in l_1; j_2 \not\in l_2$,

$$\begin{aligned} (a:1) &= S_{K,1}^{-2}S_{K,2}^{-2} E \times \\ &\quad \stackrel{a}{1} W_{k_1 i_1 j_1}(s) W_{k_1 i_1 l_1}(t) W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1}) W_{k_2 i_2 j_2}(s) W_{k_2 i_2 l_2}(t) \\ &\quad W_{k_2 i_2 j'_2}(T_{k_2 i_2 l_2}) (Y -)_{k_1 i_1 j_1} (Y -)_{k_1 i'_1 j'_1} (Y -)_{k_2 i_2 j_2} (Y -)_{k_2 i_2 j'_2} \\ &= S_{K,1}^{-2}S_{K,2}^{-2} E \times \\ &\quad \stackrel{a}{1} W_{k_1 i_1 l_1}(t) W_{k_1 i_1 j_1}(s) W_{k'_1 i'_1 j'_1}(T_{k_1 i_1 l_1}) (Y -)_{k_1 i_1 j_1} (Y -)_{k_1 i'_1 j'_1} \end{aligned}$$

$$\begin{aligned} & \mathbb{E} W_{k_2 i_2 l_2}(t) W_{k_2 i_2 j_2}(s) W_{k_2 i_2 j'_2}(T_{k_2 i_2 l_2})(Y \quad)_{k_2 i_2 j_2}(Y \quad)_{k_2 i_2 j'_2} \\ & = (\mathbb{E} R_1(s; t))^2 : \end{aligned}$$

(B.2) Let $\frac{a}{2} = f(k_1; i_1) = (k_2; i_2); (k'_1; i'_1) = (k'_2; i'_2); (k_1; i_1) \not\in (k'_1; i'_1); j_1 \not\in l_1; j_2 \not\in l_2; j'_1 \not\in j'_2 g$,

$$\begin{aligned} (a:2) &= S_{K,1}^{-2} S_{K,2}^{-2} \underset{\#}{\underset{\frac{a}{2}}{\times}} \mathbb{E} f W_{k_1 i_1 j_1}(s) W_{l_2 j_1 k_1}(s)(Y \quad)_{k_1 i_1 j_1}(Y \quad)_{l_2 j_1 k_1} g \\ &= S_{K,1}^{-2} S_{K,2}^{-2} \underset{\frac{a}{2}}{\times} (s; s) (t; t) f^2(s) f^4(t) + O_p(\quad , 2 + \quad , 2) \\ &= S_{K,1}^{-2} S_{K,2}^{-2} S_{K,3}^2 (s; s) (t; t) f^2(s) f^4(t) + O_p(\quad , 2 + \quad , 2) : \end{aligned}$$

(B.3) Let $\frac{a}{3} = f(k_1; i_1) = (k'_2; i'_2); (k_2; i_2) = (k'_1; i'_1); (k_1; i_1) \not\in (k_2; i_2); j_1 \not\in l_1; j_2 \not\in l_2 g$,

$$\begin{aligned} (a:3) &= S_{K,1}^{-2} S_{K,2}^{-2} \underset{\frac{a}{3}}{\underset{\mathfrak{h}}{\times}} \mathbb{E} W_{k_1 i_1 j_1}(s) W_{k_1 i_1 h_1}(t) W_{k_2 i_2 j'_1}(T_{k_1 i_1 h_1}) \\ &\quad W_{k_2 i_2 j_2}(s) W_{k_2 i_2 l_2}(t) W_{k_1 i_1 j'_2}(T_{k_2 i_2 l_2}) \\ &\quad (Y \quad)_{k_1 i_1 j_1}(Y \quad)_{k_2 i_2 j'_1}(Y \quad)_{k_2 i_2 j_2}(Y \quad)_{k_1 i_1 j'_2} \\ &= (\mathbb{E} R_1(s; t))^2 : \end{aligned}$$

(B.4) Let $\frac{a}{4} = f(k_1; i_1) = (k_2; i_2) = (k'_1; i'_1) = (k'_2; i'_2); j_1 \not\in l_1; j_2 \not\in l_2 g$ and omit the subscript ki ,

$$\begin{aligned} (a:4) &= S_{K,1}^{-2} S_{K,2}^{-2} \underset{\frac{a}{4}}{\underset{\cap}{\times}} \mathbb{E} W_{j_1}(s) W_{l_1}(t) W_{j'_1}(T_{l_1}) W_{j_2}(s) W_{l_2}(t) W_{j'_2}(T_{l_2}) \\ &\quad (Y \quad)_{j_1}(Y \quad)_{j'_1}(Y \quad)_{j_2}(Y \quad)_{j'_2} : \end{aligned}$$

It is the summation of the following sets,

$$\frac{a}{4:1} = f j_1 = j_2; j'_1 = j'_2; l_1 = l_2; j_1 \not\in l_1; j_2 \not\in l_2 g;$$

$$\frac{a}{4:2} = f j_1 = j_2; j'_1 \not\in j'_2; j_1 \not\in l_1; j_2 \not\in l_2 g [f j_1 \not\in j_2; j'_1 = j'_2; l_1 = l_2; j_1 \not\in l_1; j_2 \not\in l_2 g;$$

$$\frac{a}{4:3} = f j_1; j_2; j'_1; j'_2 \text{ are not equal}; j_1 \not\in l_1; j_2 \not\in l_2 g:$$

Correspondingly, we define

$$\begin{aligned}
 E_{1|4}(s; t) &= E(Y_1 | (T_1)^2 Y_2 | (T_2)^2) j T_1 = s; T_2 = t; \\
 E_{2|4}(s; t) &= E(Y_1 | (T_1)^2 Y_2 | (T_2) Y_3 | (T_3)) j T_1 = s; T_2 = t; T_3 = t; \\
 E_{3|4}(s; t) &= E(Y_1 | (T_1) Y_2 | (T_2) Y_3 | (T_3) Y_4 | (T_4)) j T_1 = s; \\
 &\quad T_2 = t; T_3 = s; T_4 = t;
 \end{aligned}$$

Then by easy calculation,

$$\begin{aligned}
 (a:4)_{\frac{a}{4}:1} &\stackrel{\dot{=}}{=} S_{K,1}^{-2} S_{K,2}^{-2} \left(\sum_{k=1}^{\infty} S_{k,3} e_{;k}^{-1} e_{;k}^{-2} R(W)^3 E_{1|4}(s; t) f(s) f^2(t) \right) ; \\
 (a:4)_{\frac{a}{4}:2} &\stackrel{\dot{=}}{=} S_{K,1}^{-2} S_{K,2}^{-2} \left(\sum_{k=1}^{\infty} S_{k,5} e_{;k}^{-1} R(W) E_{2|4}(s; t) f(s) f^4(t) \right. \\
 &\quad \left. + \sum_{k=1}^{\infty} S_{k,4} e_{;k}^{-1} e_{;k}^{-1} R(W)^2 E_{2|4}(t; s) f^2(s) f^2(t) \right) ; \\
 (a:4)_{\frac{a}{4}:3} &\stackrel{\dot{=}}{=} S_{K,1}^{-2} S_{K,2}^{-2} S_6 E_{3|4}(s; t) f^2(s) f^4(t);
 \end{aligned}$$

b. For (b) and (c), it is required that $(k_1; i_1) = (k_2; i_2) = (k'_1; i'_1)$, define

$$E_{2|3}(s; s; t) = E(Y_1 | (T_1) Y_2 | (T_2) Y_3 | (T_3)) j T_1 = s; T_2 = s; T_3 = t;$$

By easy calculation,

$$(b) + (c) \stackrel{\dot{=}}{=} S_{K,1}^{-2} S_{K,2}^{-2} (W) E_{2|3}(s; s; t) \left(\sum_{k=1}^{\infty} S_{k,3} e_{;k}^2 \right) ;$$

c. For (d), we need $(k_1; i_1) = (k_2; i_2)$, then

$$(d) \stackrel{\dot{=}}{=} \frac{1}{4} S_{K,2}^{-2} (W) f''(t) g^2(s; s) f^2(s) f^4(t) \left(\sum_{k=1}^{\infty} S_{k,4} e_{;k}^4 \right) ;$$

From the above arguments,

$$\begin{aligned}
 \text{Var } R_1 &= E(R_1^2) - (E(R_1))^2 \\
 &= O_p(S_{K,1}^{-2} S_{K,2}^{-2} S_{K,3}^2 + \sum_k S_{k,3} e_{;k}^{-1} e_{;k}^{-2} + \sum_k S_{k,4} e_{;k}^{-1} e_{;k}^{-1} + \sum_k S_{k,5} e_{;k}^{-1} \\
 &\quad + S_{K,2}^{-2} \sum_k S_{k,4} e_{;k}^4) ;
 \end{aligned}$$

Note that

$$\begin{aligned}\text{Cov}(Q_{0R_1}; \hat{R}_1) &= S_{K,1}^{-2} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{m_{ki}} m_{ki}^2 (m_{ki} - 1) \stackrel{!}{=} R(W)(s; t)f(s)f^3(t); \\ \mathbb{E}(Q_{0R_1}) &= f(t) + \frac{1}{2} (\mathbb{E}W)f''(t) \stackrel{!}{=} o_p(\cdot); \\ \text{Var}(Q_{0R_1}) &= \frac{\mathbb{E}^2}{S_{K,1}} R(W)f(t) + o_p\left(\frac{\mathbb{E}^2}{S_{K,1}}\right);\end{aligned}$$

By the delta method,

$$\begin{aligned}\text{Var}\left(\frac{\hat{R}_1}{Q_{0R_1}}\right) &= \frac{\text{Var}(\hat{R}_1)}{\mathbb{E}^2 Q_{0R_1}} + 2 \frac{\mathbb{E}(\hat{R}_1 \text{Cov}(Q_{0R_1}; \hat{R}_1))}{\mathbb{E}^3 Q_{0R_1}} + \frac{\mathbb{E}^2(\hat{R}_1 \text{Var}(Q_{0R_1}))}{\mathbb{E}^4 Q_{0R_1}} \\ &= O_p\left(S_{K,1}^{-2} S_{K,2}^{-2} S_{K,3}^2 + \sum_k s_{k,3} e_{;k}^{-1} e_{;k}^{-2} + \sum_k s_{k,4} e_{;k}^{-1} e_{;k}^{-1} + \sum_k s_{k,5} e_{;k}^{-1} \right. \\ &\quad \left. + \sum_k S_{K,2}^{-2} s_{k,4} e_{;k}^4\right);\end{aligned}$$

■

Now we give the proof of the asymptotic distribution of $e^{(K)}$ in Theorem 2.

Proof. Denote

$$f_1 = Q_{20}Q_{02} - Q_{11}^2; \quad f_2 = Q_{10}Q_{02} - Q_{01}Q_{11}; \quad f_3 = Q_{01}Q_{20} - Q_{10}Q_{11};$$

Let e be the local linear estimate based on $\hat{e}_{kijl} = (Y_{kij} - e_{kij})(Y_{kil} - e_{kil})$ which can be expressed as

$$e = \frac{f_1 R_{00} - f_2 R_{10} - f_3 R_{01}}{f_1 Q_{00} - f_2 Q_{10} - f_3 Q_{01}},$$

where for $p, q = 0, 1, 2$,

$$\begin{aligned}Q_{p,q} &= S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{m_{ki}} \sum_{1 \leq j \neq l \leq m_{ki}} W_{;k}(T_{kij} - s) W_{;k}(T_{kil} - t) \frac{T_{kij} - s}{e_{;k}}^p \frac{T_{kil} - t}{e_{;k}}^q \\ &= \sum_k !_k Q_{pq;k}; \\ R_{p,q} &= S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^{m_{ki}} \sum_{1 \leq j \neq l \leq m_{ki}} W_{;k}(T_{kij} - s) W_{;k}(T_{kil} - t) \frac{T_{kij} - s}{e_{;k}}^p \frac{T_{kil} - t}{e_{;k}}^q \hat{e}_{kijl} \\ &= \sum_k !_k R_{pq;k};\end{aligned}$$

Let

$$\begin{aligned}
e_{1;k} &= \frac{1}{e_{;k}} \frac{f_2(R_{10} - R_{00}Q_{10}=Q_{00})}{f_1Q_{10} f_2Q_{10}^2=Q_{00} f_3Q_{01}Q_{10}=Q_{00}}; \\
e_{2;k} &= \frac{1}{e_{;k}} \frac{f_3(R_{01} - R_{00}Q_{01}=Q_{00})}{f_1Q_{01} f_2Q_{01}Q_{10}=Q_{00} f_3Q_{01}^2=Q_{00}}; \\
e_{1;k}^* &= \frac{1}{e_{;k}} \frac{f_2(R_{10}^* - R_{00}^*Q_{10}=Q_{00})}{f_1Q_{10} f_2Q_{10}^2=Q_{00} f_3Q_{01}Q_{10}=Q_{00}}; \\
e_{2;k}^* &= \frac{1}{e_{;k}} \frac{f_3(R_{01}^* - R_{00}^*Q_{01}=Q_{00})}{f_1Q_{01} f_2Q_{01}Q_{10}=Q_{00} f_3Q_{01}^2=Q_{00}}; \\
R_{p;q}^* &= S_{K;2}^{-1} \underset{k=1}{\times} \underset{i=1}{\times} \underset{1 \leq j \neq l \leq m_{ki}}{\times} W_{\sim ;k}(T_{kij} - s) W_{\sim ;k}(T_{kil} - t) \frac{T_{kij}}{e_{;k}} s^p \frac{T_{kil}}{e_{;k}} t^q C_{kijl}; \\
\mathcal{E}_{pq} &= S_{K;2}^{-1} \underset{k=1}{\times} \underset{i=1}{\times} \underset{1 \leq j \neq l \leq m_{ki}}{\times} W_{\sim ;k}(T_{kij} - s) W_{\sim ;k}(T_{kil} - t) T_{kij}^p T_{kil}^q = \underset{k}{\times} !_k \mathcal{E}_{pq;k}.
\end{aligned}$$

Then we can write

$$\begin{aligned}
e &= \frac{R_{00}}{Q_{00}} \underset{k}{\times} !_k e_{1;k} \frac{\mathcal{E}_{10;k} sQ_{00}}{Q_{00}} \underset{k}{\times} !_k e_{2;k} \frac{\mathcal{E}_{01;k} tQ_{00}}{Q_{00}}; \\
&- = \frac{R_{00}}{Q_{00}} \frac{\partial}{\partial s} \frac{\mathcal{E}_{10} sQ_{00}}{Q_{00}} \frac{\partial}{\partial t} \frac{\mathcal{E}_{01} tQ_{00}}{Q_{00}}, \\
&-^* = \frac{R_{00}^*}{Q_{00}} \frac{\partial}{\partial s} \frac{\mathcal{E}_{10} sQ_{00}}{Q_{00}} \frac{\partial}{\partial t} \frac{\mathcal{E}_{01} tQ_{00}}{Q_{00}}.
\end{aligned}$$

Following similar arguments in the proof of Theorem 1, we can prove that

$$e^- = o_p 4 \min \underset{k=1}{\overset{2}{\times}} \underset{i=1}{\overset{8}{\times}} \frac{S_{K;2}^{-2}}{e_{;k}} \frac{!}{\sqrt{2}}; \quad S_{K;4} = S_{K;2}^2 \stackrel{93}{=} 5;$$

Now we derive the asymptotic distribution of $-$. We first show that $- -^* = O_p f(N_K)^{-1} g$.

(B.1) Calculate $- -^*$.

From the definition of $-$ and $-^*$, we have $- -^* = R_{00} = Q_{00}$. Note that

$$\begin{aligned}
R_{00} &= S_{K;2}^{-1} \underset{k=1}{\times} \underset{i=1}{\times} \underset{j \neq l}{\times} W_{kij}(s) W_{kil}(t) \mathcal{E}_{kijl} \\
&= S_{K;2}^{-1} \underset{k=1}{\times} \underset{i=1}{\times} \underset{j \neq l}{\times} W_{kij}(s) W_{kil}(t) f(Y_{kij} - e_{kij})(Y_{kil} - e_{kil}) + (Y_{kij} - e_{kij})(Y_{kil} - e_{kil}) \\
&\quad + (e_{kij} - e_{kil})(Y_{kil} - e_{kil}) + (e_{kij} - e_{kil})(e_{kil} - e_{kil})g
\end{aligned}$$

$$\begin{aligned}
& \stackrel{\cdot}{=} R_{00}^* + S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^K \sum_{j \neq i} W_{kij}(s) W_{kil}(t) f(Y_{kij} - e_{kij}) (-e_{kil} - e_{kil}) \\
& \quad + (-e_{kil} - e_{kil})(Y_{kil} - e_{kil}) g;
\end{aligned}$$

Recall that in S.1, we have

$$e(t) - \bar{e}(t) = o_p(\bar{e}(t) - e(t));$$

where

$$\begin{aligned}
\bar{e}(t) &= \frac{R_0}{Q_0} \frac{G_1}{Q_0} '(t) + t \frac{G_0}{Q_0} '(t); \\
Q_r(t) &= \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^K \sum_{j=1}^K W_{\sim, k}(T_{kij} - t) \frac{T_{kij} - t}{e_{,k}}^r; \\
R_r(t) &= \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^K \sum_{j=1}^K W_{\sim, k}(T_{kij} - t) \frac{T_{kij} - t}{e_{,k}}^r Y_{kij}; \\
G_r(t) &= \frac{1}{S_{K,1}} \sum_{k=1}^K \sum_{i=1}^K \sum_{j=1}^K W_{\sim, k}(T_{kij} - t) T_{kij}^r;
\end{aligned}$$

then one can obtain

$$\begin{aligned}
& S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^K \sum_{j \neq i} W_{kij}(s) W_{kil}(t) (Y_{kij} - e_{kij}) (-e_{kil} - e_{kil}) \\
& \stackrel{\cdot}{=} S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^K \sum_{j \neq i} W_{kij}(s) W_{kil}(t) (Y_{kij} - e_{kij}) (-e_{kil} - \bar{e}_{kil}) \\
& = S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^K \sum_{j \neq i} W_{kij}(s) W_{kil}(t) (Y_{kij} - e_{kij}) (-e_{kil} - (T_{kil}) - \frac{R_0}{Q_0} + \frac{G_1}{Q_0} '(T_{kil}) - T_{kil} \frac{G_0}{Q_0} '(T_{kil})) \\
& = S_{K,2}^{-1} Q_0^{-1} \sum_{k=1}^K \sum_{i=1}^K \sum_{j \neq i} W_{kij}(s) W_{kil}(t) (Y_{kij} - e_{kij}) \\
& \quad Q_0(T_{kil}) (T_{kil}) - R_0(T_{kil}) + G_1(T_{kil}) '(T_{kil}) - T_{kil} G_0(T_{kil}) '(T_{kil}) \\
& = S_{K,1}^{-1} S_{K,2}^{-1} Q_0^{-1} \sum_{k=1}^K \sum_{i=1}^K \sum_{j \neq i} \sum_{k'=1}^{K'} \sum_{i'=1}^{i'} \sum_{j'=1}^{j'} W_{\sim, k'}(T_{kij} - s) W_{\sim, k'}(T_{kil} - t) W_{\sim, k'}(T_{k'ij'} - T_{kil}) \\
& \quad (Y_{kij} - (T_{kij})) f(T_{kil}) - Y_{k'ij'} + T_{k'ij'} '(T_{kil}) - T_{kil} '(T_{kil}) g;
\end{aligned}$$

Take the Taylor expansion of (T_{kil}) at $(T_{k'ij'})$ in the last bracket,

$$S_{K,2}^{-1} \sum_{k=1}^K \sum_{i=1}^K \sum_{j \neq i} W_{kij}(s) W_{kil}(t) (Y_{kij} - e_{kij}) (-e_{kil} - e_{kil})$$

$$\begin{aligned}
& \stackrel{\doteq}{=} S_{K,1}^{-1} S_{K,2}^{-1} Q_0^{-1} \underset{k=1}{\times} \underset{i=1}{\times} \underset{j \neq i}{\times} \underset{k'=1}{\times} \underset{i'=1}{\times} \underset{j'=1}{\times} \underset{n}{\underset{\cap}{\times}} W_{\sim ;k'}(T_{kij} - s) W_{\sim ;k'}(T_{kil} - t) \\
& \quad W_{\sim ;k'}(T_{k'ij'}) - T_{kil})(Y_{kij} - (T_{kij})) Y_{k'ij'} - (T_{k'ij'}) + \\
& \quad '(T_{kil}) - '(T_{k'ij'}) (T_{k'ij'} - T_{kil}) + \frac{1}{2} ''(T_{kil})(T_{k'ij'})^2 \circ :
\end{aligned}$$

Then take the Taylor expansion of $'(T_{k'ij'})$ at $'(T_{kil})$ in the last bracket,

$$\begin{aligned}
& S_{K,2}^{-1} \underset{k=1}{\times} \underset{i=1}{\times} \underset{j \neq i}{\times} W_{kij}(s) W_{kil}(t) (Y_{kij} - (T_{kij})) (Y_{kil} - e_{kil}) \\
& \stackrel{\doteq}{=} S_{K,1}^{-1} S_{K,2}^{-1} Q_0^{-1} \underset{k=1}{\times} \underset{i=1}{\times} \underset{j \neq i}{\times} \underset{k'=1}{\times} \underset{i'=1}{\times} \underset{j'=1}{\times} \underset{n}{\underset{\cap}{\times}} W_{\sim ;k'}(T_{kij} - s) W_{\sim ;k'}(T_{kil} - t) \\
& \quad W_{\sim ;k'}(T_{k'ij'}) - T_{kil})(Y_{kij} - (T_{kij})) Y_{k'ij'} - (T_{k'ij'}) - \frac{1}{2} ''(T_{kil})(T_{k'ij'})^2 \circ :
\end{aligned}$$

Denote

$$\begin{aligned}
\Delta R_1 & = S_{K,1}^{-1} S_{K,2}^{-1} \underset{k=1}{\times} \underset{i=1}{\times} \underset{j \neq i}{\times} \underset{k'=1}{\times} \underset{i'=1}{\times} \underset{j'=1}{\times} W_{\sim ;k'}(T_{kij} - s) W_{\sim ;k'}(T_{kil} - t) \\
& \quad W_{\sim ;k'}(T_{k'ij'}) - T_{kil})(Y_{kij} - (T_{kij})) Y_{k'ij'} - (T_{k'ij'}) - \frac{1}{2} ''(T_{kil})(T_{k'ij'})^2 ; \\
\Delta R_2 & = S_{K,1}^{-1} S_{K,2}^{-1} \underset{k=1}{\times} \underset{i=1}{\times} \underset{j \neq i}{\times} \underset{k'=1}{\times} \underset{i'=1}{\times} \underset{j'=1}{\times} W_{\sim ;k'}(T_{kij} - s) W_{\sim ;k'}(T_{kil} - t) \\
& \quad W_{\sim ;k'}(T_{k'ij'}) - T_{kil})(Y_{kil} - (T_{kil})) Y_{k'ij'} - (T_{k'ij'}) - \frac{1}{2} ''(T_{kil})(T_{k'ij'})^2 ;
\end{aligned}$$

$$Q_{0R1} = Q_0(T_{kil}); \quad Q_{0R2} = Q_0(T_{kij});$$

It can be written as

$$R_{00} \stackrel{\doteq}{=} R_{00}^* + \frac{R_1}{Q_{0R1}} + \frac{R_2}{Q_{0R2}};$$

The last terms in the right hand side are of the same order. By Lemma S.1, we have

$$\begin{aligned}
E(\Delta R_1 - Q_{0R1}) & = O_p(S_{K,1}^{-1} S_{K,2}^{-1} S_{K,3}); \\
\text{Var}(\Delta R_1 - Q_{0R1}) & = O_p(S_{K,1}^{-2} S_{K,2}^{-2} S_{K,3}^2 (1 + S_{K,1}^{-1}) + S_{K,3}^{-1} + S_{K,4}^{-1} + S_{K,5}^{-1} + S_{K,6}^{-1} + S_{K,2}^{-2} S_4 (\cdot)^2) ;
\end{aligned}$$

Hence

$$\begin{aligned}
& - - * = O_p(S_{K,1}^{-1} S_{K,2}^{-1} S_{K,3}^2 (1 + S_{K,1}^{-1}) + S_{K,3}^{-1} + S_{K,4}^{-1} + S_{K,5}^{-1} + S_{K,6}^{-1} + S_{K,2}^{-2} S_4^{\frac{1}{2}} (\cdot)^2) ;
\end{aligned}$$

(B.2) Asymptotic distribution of $-\ast$.

Define

$$-\ast = \frac{R_{00}^*}{Q_{00}} - \frac{\partial}{\partial s} \frac{\mathbb{E}_{10}}{Q_{00}} s Q_{00} - \frac{\partial}{\partial t} \frac{\mathbb{E}_{01}}{Q_{00}} t Q_{00}.$$

The asymptotic normality of $-\ast$ is accomplished by the asymptotic joint normality of $(R_{00}^*, \mathbb{E} R_{00}^*, \mathbb{E} \mathbb{E}_{00}, \mathbb{E} \mathbb{E}_{10}, \mathbb{E} \mathbb{E}_{10}, \mathbb{E} \mathbb{E}_{01})$. Specifically, let

$$\mathbb{E}_1(s; t) = \mathbb{E} f(Y_1 - (T_1))^2 (Y_2 - (T_2))^2 | T_1 = s; T_2 = t g;$$

$$\mathbb{E}_2(s; t) = \mathbb{E} f(Y_1 - (T_1))^2 (Y_2 - (T_2)) (Y_3 - (T_3)) | T_1 = s; T_2 = t; T_3 = t g;$$

By easy calculation, we have

$$\begin{aligned} \mathbb{E} \mathbb{E}_{pq} &= s^p t^q f(s) f(t) + \frac{1}{2} (\mathcal{W}) f s^p t^q f''(s) f(t) + s^p t^q f(s) f'(t) g^{-2} \\ &\quad + \frac{1}{2} (\mathcal{W}) f 2 p f'(s) f(t) + 2 q f'(t) f(s) g^{-2} + o_p^{-2}; \\ \mathbb{E} R_{00}^* &= (s; t) f(s) f(t) + \frac{1}{2} (\mathcal{W})^{-2} \frac{\partial^2}{\partial s^2} (s; t) f(s) f(t) \\ &\quad + 2 \frac{\partial}{\partial s} (s; t) f'(s) f(t) + (s; t) f''(s) f(t) + \frac{\partial^2}{\partial t^2} (s; t) f(t) f(s) \\ &\quad + 2 \frac{\partial}{\partial s} (s; t) f'(t) f(s) + (s; t) f''(t) f(s) + o_p^{-2}; \\ \text{Var}(R_{00}^*) &= f^2 + I(s = t) g \mathcal{R}(\mathcal{W})^2 E_1(s; t) f(s) f(t) S_{K,2}^{-2} \\ &\quad + R(\mathcal{W}) f^2(s) f(t) E_2(t; s) + f(s) f(t)^2 E_2(s; t) S_{K,2}^{-1} \underset{k=1}{\overset{\#}{\times}} \frac{S_{k,3}}{\mathbb{E}_{:,k}} \\ &\quad + V_3(s; t) f^2(s) f^2(t) S_{K,2}^{-2} S_4 + o_p S_{K,2}^{-2} \\ &\quad + o_p S_{K,2}^{-2} \underset{k=1}{\overset{\#}{\times}} \frac{S_{k,3}}{\mathbb{E}_{:,k}} + o_p S_{K,2}^{-2} S_4; \end{aligned}$$

$$\begin{aligned} \text{Cov}(\mathbb{E}_{pq}, \mathbb{E}_{p'q'}) &= f^2 + I(s = t) g s^{p+p'} t^{q+q'} \mathcal{R}(\mathcal{W})^2 f(s) f(t) S_{K,2}^{-2} \\ &\quad + R(\mathcal{W}) f f^2(s) f(t) + f(s) f^2(t) g S_{K,2}^{-2} \underset{k=1}{\overset{\#}{\times}} \frac{S_{k,3}}{\mathbb{E}_{:,k}} \\ &\quad + o_p S_{K,2}^{-2} + o_p S_{K,2}^{-2} \underset{k=1}{\overset{\#}{\times}} \frac{S_{k,3}}{\mathbb{E}_{:,k}}; \\ \text{Cov}(\mathbb{E}_{pq}, R_{00}^*) &= f^2 + I(s = t) g s^p t^q \mathcal{R}(\mathcal{W})^2 (s; t) f(s) f(t) S_{K,2}^{-2} \end{aligned}$$

$$\begin{aligned}
& + R(W) (s; t) f f^2(s) f(t) + f(s) f^2(t) g S_{K,2}^{-2} \underset{k=1}{\overset{\#}{\times}} \frac{s_{k,3}}{e_{,k}} \\
& + o_p S_{K,2}^{-2} \underset{k=1}{\overset{\#}{\times}} \frac{s_{k,3}}{e_{,k}} ;
\end{aligned}$$

From the delta method, $\bar{-(t)}$ follows the asymptotic normality that

$$\bar{-(s,t)} = (s,t) - \frac{1}{2} \bar{(W)} \left(\frac{\partial^2}{\partial s^2}(s,t) + \frac{\partial^2}{\partial t^2}(s,t) \right) + o_p \bar{(-2)} \stackrel{d}{\rightarrow} N(0,1);$$

where

$$\begin{aligned}
\bar{-(s,t)} &= f^2(s,t) + I(s=t) g S_{K,2}^{-1} \bar{(-2)} R(W)^2 \frac{V_1(s,t)}{f(s)f(t)} \\
&+ S_{K,2}^{-2} \underset{k=1}{\overset{\#}{\times}} \frac{s_{k,3}}{e_{,k}} R(W) \frac{f(s)V_2(t,s) + f(t)V_2(s,t)}{f(s)f(t)} + V_3(s,t) S_4 = S_{K,2}^2;
\end{aligned}$$

The proof is completed combining the above arguments. ■

S.3. PROOF OF THEOREM 3

We mention that Theorem 3 and Theorem 4 can be extended to the standard d -dimensional local linear regression. To see this, we introduce the following notations and assumptions.

Notations and assumptions for d -dimensional local linear regression. The underlying model is $Y = m(X) + \epsilon$ where $X \in \mathbb{R}^d$, $E\epsilon = 0$ and $\text{Var } \epsilon = \Sigma(X)$. Suppose we observe $f(Y_{ki}; X_{ki} : i = 1, \dots, n_k)g$ in the k th block and $N_K = \sum_{k=1}^K n_k$, i.e., we use n_K to denote the sub-sample size of the K th data block and $N_K = \sum_{k=1}^K n_k$ to denote the full sample size up to time K , which correspond to $s_{K,1}; S_{K,1}$ in (6) for mean estimation and $s_{K,2}; S_{K,2}$ for covariance estimation, respectively. We impose the following assumptions which are in parallel to (A.1)–(A.3) and (A.6).

- (B.1) Observations $fX_{ki} : i = 1, \dots, n_k; k = 1, \dots, K$ are i.i.d. copies of a random variable X defined on $[0, 1]^d$ whose density $f(\cdot)$ is bounded away from 0 with bounded second derivative.
- (B.2) The second and fourth derivatives of regression function $m(x)$ are bounded and continuous.

(B.3) Noises ε_{ki} are i.i.d. with mean 0 and variance $\sigma^2(x)$.

(B.4) The block size $Kn_k = N_K / 1$ as $K \rightarrow 1$ for $k = 1; 2; \dots; K$.

The optimal bandwidth at time K is $h_*^{(K)} = N_K^{-1/(4+d)}$. The candidate sequence is $f_i^{(K)} g_{i=1}^L$ and the pseudo-bandwidths are $f\bar{e}_k^{(K)} : k = 1; \dots; Kg$ for $j = 1; \dots; d$. The centroids for the candidate sequence are $f_i^{(K)} : i = 1; \dots; Lg$. Further define

$$f_i^{(K)} = \frac{\sum_{k=1}^K n_k}{N_K} e_k^{(K)} - i;$$

We also mention that the mean and covariance estimates corresponds the cases $d = 1$ and $d = 2$, respectively.

When estimating μ , all observations are pooled together and the index of subjects become invalid, which is intrinsically one-dimensional local linear regression. Hence one can derive the following result for standard d -dimensional local linear regression by the same arguments in Theorem 1.

Lemma S. 2. Suppose assumptions (B.1)–(B.3) and (A.8)–(A.9) hold. If

$$\hat{h}_*^{(K)} = h_*^{(K)} = o_p(N_K^{-1/(4+d)}) \quad ; \quad j = 1; \dots; d; \quad (\text{S.3})$$

then a fixed interior point $x \in (0; 1)^d$, as $K \rightarrow 1$, the online estimate $\hat{m}^{(K)}(x)$ satisfies

$$N_K \hat{h}_*^{(K)} \frac{O_1}{(-d)^2} 4 \hat{m}^{(K)}(x) - m(x) - \frac{1}{2} \langle W \rangle_{j=1}^8 = \frac{m''_j(x)}{2} + o_p(\sqrt{2})^5 \notin N(0; \frac{R(W)^{d-2}(x)}{f(x)});$$

We prove in Section S.3 that (S.3) is satisfied by the proposed online method.

Denote $X = fX_{ki} \in \mathbb{R}^d : k = 1; \dots; K; i = 1; \dots; n_k g_i = \int_{j=1}^d \partial^2 m(x) = \partial x_j^2 g f(x) dx$ and $\hat{R} = \int_{[0,1]^d} R(W)^{d-2}(x) dx$. Similarly, we adopt another online local cubic and online local linear to estimate μ and σ^2 , which are denoted as $\hat{e}^{(K)}$ and $\hat{e}^{(K)}$, respectively. Then the optimal bandwidth and the online estimated bandwidth for $m(\cdot)$ are

$$h_*^{(K)} = \frac{1}{\sigma^2(W)} N_K^{-1/(4+d)}, \quad \hat{h}_*^{(K)} = \frac{\hat{e}^{(K)}}{\sigma^2(W)^{1/(4+d)}} N_K^{-1/(4+d)}; \quad (\text{S.4})$$

The candidate sequence for μ and σ^2 are $f_i^{(K)} g_{i=1}^L$ and $f_i^{(K)} g_{i=1}^L$, respectively, and the pseudo-bandwidths are $f\bar{e}_{;k}^{(K)} : k = 1; \dots; Kg$ and $f\bar{e}_{;k}^{(K)} : k = 1; \dots; Kg$. Then we have the following conclusion.

Lemma S. 3. Suppose assumptions (B.1)–(B.4) and (A.8)–(A.9) hold. When the online bandwidths and candidate sequence for \hat{h} and \hat{h}_* are

$$h^{(K)} = GN_K^{-1/(6+d)}; \quad h_*^{(K)} = RN_K^{-1/(4+d)}; \quad 1 < G, R < 1;$$

$$\hat{h}_{;J}^{(K)} < \dots < \hat{h}_{;1}^{(K)} = h^{(K)}; \quad \hat{h}_{;J}^{(K)} < \dots < \hat{h}_{;1}^{(K)} = h_*^{(K)};$$

then as $K \rightarrow \infty$, $\hat{h}^{(K)}$ in (S.4) satisfies

$$\hat{h}^{(K)} = h_*^{(K)}$$

and then $e^{(K)} = O_p(S_{K,1}^{-2=5})$. For , one can prove that the convergence rate of $e^{(K)}$ is dominated by $E_K(s)(t)$ in step 3 of case (b) in Appendix B which equals $O_p(S_{K,2}^{-1=3})$. For dense data, we need only check the order of e_K^2 . From Zhang and Chen (2007), " \hat{e}_{kij} " in step 2 of (c) in Appendix B is of order $(S_{K,1})^{-2=25}$, recall that \hat{b}_{kij}^2 is a debiased version of " \hat{e}_{kij} " and e_K^2 is the average of \hat{b}_{kij}^2 , then $e_K^2 = O_p(S_{K,1}^{-12=25})$. In all the cases, estimates of are dominates those of , which completes the proof of Theorem 3.

S.4. PROOF OF THEOREM 4

Note that the asymptotic distributions of e and e have the consistent form with the standard d -dimensional local linear regression, i.e., the bias is of $O_p(h^2)$ and the variance is of $O_p(N^{-1}h^{-d})$, where $d = 1$ for mean estimation and $d = 2$ for covariance estimation. The lower bound of the relative efficiency also holds for the standard d -dimensional local linear regression. We adopt the notations of the d -dimensional local linear in the proof for notation conciseness.

Proof. Recall that $\hat{h}^{(K)} = \frac{(K)}{1} > \dots > \frac{(K)}{L}$, without loss of generality, we assume $\frac{(K)}{l} = g(l)\hat{h}^{(K)}$ where $g(l) = 1, l = 1, \dots, L$. Note that $\hat{h}^{(K)} = o_p(N_K^{-1=(d+4)})$ and the candidate bandwidth sequence satisfies

$$\frac{e_k^{(K)}}{\hat{h}^{(K)}} = g(l) \frac{N_k}{N_K}^{-\frac{1}{d+2}} + o_p(1);$$

Write $g(l) = fg(l)^{1=}$ and note that $N_k=N_K$ grows linearly with respect to k , then shall be $1=(d+2)$ and the optimal $g(l)^{1=}$ shall be linear between $(0;1)$. Hence, the optimal , is as in (19).

Next we derive the asymptotic lower bound for the relative efficiency of our local linear estimator. Writing $\hat{h}^{(K)} = h_*^{(K)}f1 + O_p(N_K^{-2=(d+6)})g$, one can derive

$$eff(\hat{h})^{-1} = \frac{d}{d+4} \sum_{k=1}^8 \frac{n_k}{N_K} \frac{e_k^{(K)}}{\hat{h}^{(K)}} !_2^2 = + \frac{4}{d+4} \sum_{k=1}^8 \frac{n_k}{N_K} \frac{e_k^{(K)}}{\hat{h}^{(K)}} !_{-1}^2 = + O_p(N_K^{-\frac{2}{6+d}}); \quad (S.5)$$

As shown in Figure 1, when K tends large, there exists a breakpoint K_0 : when $k < K_0$, e_k equals the last candidate $(1=L)^{1=(d+4)}\hat{h}_k$ for the limit number of candidates and for $k > K_0$,

there are sufficient and close candidates to make a choice. Hence

$$\frac{e_k^{(K)}}{\theta^{(K)}} = \begin{cases} \underset{8}{\gtrless} L^{-\frac{1}{d+4}} \frac{N_k}{N_K}^{-\frac{1}{d+4}} + O_p(N_k^{-\frac{2}{6+d}}) & ; \quad k \leq K_0 \\ \underset{n}{\gtrless} 1 + \frac{N_k}{N_K} \frac{I_k}{L} O^{\frac{1}{d+4}} + O_p(N_k^{-\frac{2}{6+d}}) & ; \quad k > K_0 \end{cases}$$

where $I_k = \operatorname{argmin}_{1 \leq l \leq L} j N_k = N_K - I = Lj$. Then when $k \leq K_0$,

$$\begin{aligned} \underset{k=1}{\overset{K_0}{\times}} \frac{n_k}{N_K} \left| \frac{e_k^{(K)}}{\theta^{(K)}} \right|^2 &= \frac{d+4}{d+2} L^{-\frac{2}{d+4}} \frac{N_{K_0}}{N_K}^{1-\frac{2}{d+4}} + O_p(N_k^{-\frac{2}{6+d}}) ; \\ \underset{k=1}{\overset{K_0}{\times}} \frac{n_k}{N_K} \left| \frac{e_k^{(K)}}{\theta^{(K)}} \right|^{-d} &= \frac{d+4}{2d+4} L^{\frac{d}{d+4}} \frac{N_{K_0}}{N_K}^{1+\frac{d}{d+4}} + O_p(N_k^{-\frac{2}{6+d}}) ; \end{aligned}$$

and when $k > K_0$,

$$\begin{aligned} \underset{k=K_0+1}{\overset{K}{\times}} \frac{n_k}{N_K} \left| \frac{e_k^{(K)}}{\theta^{(K)}} \right|^2 &= 1 \frac{N_{K_0}}{N_K} + O_p(N_k^{-\frac{2}{6+d}}) ; \\ \underset{k=K_0+1}{\overset{K}{\times}} \frac{n_k}{N_K} \left| \frac{e_k^{(K)}}{\theta^{(K)}} \right|^{-1} &= 1 \frac{N_{K_0}}{N_K} + O_p(N_k^{-\frac{2}{6+d}}) ; \end{aligned}$$

The property of breakpoint K_0 also guarantees that

$$\frac{1}{L} \underset{k=1}{\overset{K_0}{\times}} \frac{n_k}{N_{K_0}} \theta^{(k)} = \theta^{(K)} ;$$

which is equivalent to

$$\frac{1}{L} \underset{k=1}{\overset{K_0}{\times}} \frac{n_k}{N_{K_0}} \frac{\theta^{(k)}}{\theta^{(K)}} \frac{h_*^{(k)}}{h_*^{(K)}} \frac{h_*^{(K)}}{h_*^{(K)}} = \frac{1}{L} \underset{k=1}{\overset{K_0}{\times}} \frac{n_k}{N_{K_0}} f 1 + O_p(N_k^{-\frac{2}{6+d}}) g \frac{N_k}{N_K}^{\frac{1}{d+4}} = 1 ;$$

and under Assumption (A.5), we obtain

$$\frac{N_{K_0}}{N_K} = \frac{d+4}{d+3} \frac{d+4}{d+2} \frac{1}{L} + O_p(N_k^{-\frac{2}{6+d}}) ;$$

Denote $\gamma = N_{K_0}/N_K$, one can derive from (S.5) that

$$\begin{aligned} \operatorname{eff}(\theta)^{-1} &= \frac{d}{d+4} \frac{d+4}{d+2} L^{-\frac{2}{d+4} - 1 - \frac{2}{d+4}} + 1 \\ &\quad + \frac{4}{d+4} \frac{d+4}{2d+4} L^{\frac{d}{d+4} - 1 + \frac{d}{d+4}} + 1 + O_p(N_k^{-\frac{2}{6+d}}) ; \end{aligned} \quad (\text{S.6})$$

Note that (S.6) is strictly increasing with respect to λ . Hence we have

$$eff(\lambda)^{-1} = 1 + \frac{2c_1 d + 4c_2}{d+4} L^{-1} + \frac{d}{d+4} c_1^2 L^{-2},$$

where

$$c_1 = \frac{(d+4)^{d+3}}{(d+3)^{d+2}(d+2)} \quad ; \quad c_2 = \frac{(d+4)^{2d+5}}{(d+3)^{2d+4}(2d+4)} \quad ; \quad \frac{d+4}{d+3} \quad ;$$

This completes the proof of Theorem 4. ■

REFERENCES

- Fan, J. and Yao, Q. (1998), 'Efficient estimation of conditional variance functions in stochastic regression', *Biometrika* **85**(3), 645{660.
- Wand, M. P. and Jones, M. C. (1994), *Kernel Smoothing*, Chapman and Hall.
- Zhang, J. and Chen, J. (2007), 'Statistical inferences for functional data', *The Annals of Statistics* **35**(3), 1052{1079.
- Zhang, X. and Wang, J.-L. (2016), 'From sparse to dense functional data and beyond', *The Annals of Statistics* **44**(5), 2281{2321.